## Network Modeling

#### Viviana Amati Jürgen Lerner Bobo Nick

Dept. Computer & Information Science University of Konstanz

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## Outline

#### Introduction

Where are we going?

#### The Stochastic actor-oriented model

Data and model definition Model specification Simulating network evolution Parameter estimation: MoM and MLE Parameter interpretation

Extending the model: analyzing the co-evolution of networks and behavior

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Motivation Selection and influence Model definition and specification Simulating the co-evolution of networks and behavior Parameter estimation Parameter interpretation

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Model	Main feature	Real data
9( <i>n</i> , <i>p</i> )	ties are independent	ties are usually dependent



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$\mathfrak{G}(n,p)$	ties are independent	ties are usually dependent
Preferential attachment	based on degree distribution	there are other structural properties

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ERGM	class of models	reasonable representation of the data

These are models for cross-sectional data









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Network are dynamic by nature. How to model network evolution?

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Network are dynamic by nature. How to model network evolution?

We need a model for longitudinal data

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The *Teenage Friends and Lifestyle Study* analyzes smoking behavior and friendship

Data (available from www.stats.ox.ac.uk/ $\sim$  snijders/siena/)

- One school year group from a Scottish secondary school starting at age 12-13 years, was monitored over 3 years;
- Data were collected using questionnaires at approximately one year interval
  - 1. Friendship relation: each pupil can name up to six friend
  - 2. Individual information and lifestyle elements: gender, age, substance use, smoking of parents and siblings etc.

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The dataset is depicted by the following graphs where:

arrows denote friendship relation

gender is represented by the shape of nodes (girls: circles, boys: squares)

smoking behavior is depicted by the color of nodes (non: blue, occasional: gray, regular: black)



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- Is there any tendency in friendship formation towards reciprocity?







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- Is there any tendency in friendship formation towards transitivity?



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- Is there any homophily in friendship formation with respect to gender?







- Is there any homophily in friendship formation with respect to gender?





- Is there any homophily in friendship formation with respect to smoking behavior?



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# Solution



Stochastic actor-oriented model (SAOM)

#### Aim

Explain network evolution as a result of

- endogenous variables: structural effects (e.g. reciprocity, transitivity, etc.)
- exogenous variables: actor and dyadic covariates (gender, age, ethnicity, etc.)

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simultaneously

Definition

Let  $(\Omega, P)$  be a probability space. A (real-valued) random variable (r.v.) is a function  $X : \Omega \to \mathbb{R}$ .

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#### Example





Random experiment

#### Definition

A random variable X is called *(absolutely) continuous* if there exists a function  $f_X(x) : \mathbb{R} \to \mathbb{R}^+$  such that

$$F_X(x) = P(X \le x) = \int_{-\infty}^{x} f_X(u) du \quad \forall x \in \mathbb{R}$$

 $F_X(x)$  is the cumulative distribution function (c.d.f)

$$f_X(x)$$
 is the probability density function (p.d.f)  
-  $f_X(x) \ge 0 \quad \forall x \in \mathbb{R}$   
-  $P(X \in \mathbb{R}) = \int_{-\infty}^{+\infty} f_X(x) dx = 1$ 

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Be careful about the word *continuous*!!!

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The p.d.f.  $f_X(x)$  allows to compute all the probability statements about X. For instance, the probability that X takes values in [a, b] is

$$P(a \le X \le b) = \int_a^b f_X(x) dx$$

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#### Geometrical interpretation



Intuition suggests that

$$P(X=x) = \int_{x}^{x} f_X(u) du = 0$$

Thus, we cannot determine a continuous random variable via its "mass function"

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## Background: Exponential random variable

Definition

A continuous random variable X whose p.d.f. is given by

$$f_X(x) = \begin{cases} \lambda e^{-\lambda x} & \text{if } x \ge 0 \\ 0 & \text{otherwise} \end{cases}$$

is said to be an *Exponential* random variable with rate  $\lambda > 0$ .



## Background: Exponential random variable

The c.d.f. of X is

$$F_X(x) = \left\{ egin{array}{ccc} 1 - e^{-\lambda x} & ext{if} & x \geq 0 \ 0 & ext{otherwise} \end{array} 
ight.$$

In fact

$$F_X(x) = P(X \le x) = \int_{-\infty}^{+\infty} f_X(x) dx = \int_0^x \lambda e^{-\lambda x} dx = 1 - e^{-\lambda x}$$

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## Background: Exponential random variable

The Exponential r.v. has an important property: the memoryless property

Definition

A r.v. X is memoryless if

$$P(X > s + t | X > t) = P(X > s) \quad \forall s, t > 0$$

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It is easy to prove the memoryless property for the Exponential r.v.

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It is easy to prove the memoryless property for the Exponential r.v. Proof.

$$P(X > s + t | X > t) = \frac{P(X > t + s \cap X > t)}{P(X > t)} = \frac{P(X > t + s)}{P(X > t)} = \frac{1 - P(X \le t + s)}{1 - P(X \le t)} = \frac{1 - 1 + e^{-\lambda(t + s)}}{1 - 1 + e^{-\lambda t}} = e^{-\lambda s} = P(X > s)$$

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## Definition

A stochastic process is a collection  $\{X(t), t \in T\}$  of random variables  $X(t) : \Omega \longmapsto \mathbb{R}$ 

 $\forall t \in T \mapsto X(t) : \Omega \to \mathbb{R}$ 

- T = index set (usually interpreted as time)
- ${\boldsymbol{\mathbb{S}}} = {\sf state} \; {\sf space}$

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Example

X(t) = the outcome of flipping a coin

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Example X(t) = the outcome of flipping a coin  $S = \{-1, 1\}$ , where -1 =tail 1 =head  $T = \{1, 2, \dots\}$ 

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Example X(t) = the outcome of flipping a coin  $S = \{-1, 1\}$ , where -1 = tail 1 = head $T = \{1, 2, \dots\}$ X(t) 1 -٠ 0 2 3 4 5 6 7 8 9 1 -1 =

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 $\{X(t), t \in T\}$  is a discrete-time stochastic process with a discrete (or finite) state space

Example

X(t) = the number of telephone call at a switchboard of a company from 8 a.m. to 8 p.m.

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$$S = \{0, 1, 2, \cdots\}$$
  
 $T = [0, 12]$ 

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X(t) = the temperature in a room at each time moment

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## Example

X(t) = the temperature in a room at each time moment

# $\begin{array}{l} \mathbb{S} = [18, 25] \\ \mathbb{T} = [0, 24] \end{array}$

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 $\{X(t), t \in T\}$  is a continuous-time stochastic process with a continuous state space

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Different stochastic process can be defined according to the nature of the state space S, the index set T and the dependence relations existing among the random variables X(t).

We will consider **continuous-time Markov chains continuous-time** = the process evolves in continuous time

**Markov** = the process have the Markov property

chains = the state space is a finite set

#### Definition

 $\{X(t), t \in T\}$  has the *Markov property* if for any state  $x \in S$ , and for any pair of time points  $t_i < t_i$ 

 $P(X(t_j) = x(t_j)|X(t) = x(t) \text{ for all } t \leq t_i) = P(X(t_j) = x(t_j)|X(t_i) = x(t_i))$ 

The future depends on the past and on the present only through the present

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#### Definition

A continuous-time Markov chain  $\{X_t, t \ge 0\}$  is a finite state, continuous-time stochastic process having the Markovian property

## Example

X(t) = # of goals that a given soccer player scores by time t (time played in official matches)

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- 2. the time is continuous: time played in official games
- 3. the process  $\{X(t), t \ge 0\}$  has the Markov property.



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#### Holding times and the jump chain

1. Holding time: for each state *i*, the amount of time we spend in that state is an exponentially distributed random variable, with parameter  $\lambda_i$ 

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#### Holding times and the jump chain

- 1. Holding time: for each state *i*, the amount of time we spend in that state is an exponentially distributed random variable, with parameter  $\lambda_i$
- 2. Jump chain: is described by a jump matrix  $P = (p_{ij} : i, j \in S)$  which satisfies the following properties:

$$p_{ij} \geq 0$$
  $\sum_{j \in S} p_{ij} = 1$   $orall i, j \in S$ 

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P describes the probability of going to state j when we make a jump out a state i.

 $p_{ij} = P(X(t') = j | X(t) = i$ , given the opportunity to leave state i), t' > t



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## Example

$$P = \left[ \begin{array}{rrrrr} 0.1 & 0 & 0.6 & 0.3 \\ 0.8 & 0.1 & 0.1 & 0 \\ 0.05 & 0.5 & 0.05 & 0.4 \\ 0.6 & 0.1 & 0.15 & 0.15 \end{array} \right]$$

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#### The rate matrix

The rate matrix  $Q = (q_{ij} : i, j \in S)$  defines the rate at which the process jump from the state *i* to the state *j* in a short time interval.

It satisfies the following statements:

1. 
$$0 < -q_{ii} < \infty \quad \forall i \in S$$

2. 
$$q_{ij} > 0 \quad \forall i \neq j, i, j \in S$$

3. 
$$\sum_{j\in\mathbb{S}}q_{ij}=0 \quad \forall i\in S$$

The generic entry  $q_{ij}$  of this matrix gives the rate of transition from state *i* to state *j* is strictly related to the weighting times and the jump matrix. In particular:

$$q_{ij} = \begin{cases} \lambda_i p_{ij} & \text{if } j \neq i \\ \\ -\lambda_i & \text{if } j = i \end{cases}$$

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Background: adjacency matrix and directed relations

Social network consists of a set of actors  ${\mathcal N}$  and the relation  ${\mathcal R}$  existing among them



Adjacency matrix=X

-	0	0	0	0
1	-	1	0	0
0	0	-	0	0
0	1	1	-	0
1	1	0	0	-

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In the following network will be represented as an adjacency matrix X

Background: adjacency matrix and directed relations

Adjacency matrix X (or digraph):

- n = # of actors

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- the actor index the column and the rows of the  $(n \times n)$  matrix X

$$x_{ij} = \begin{cases} 1 & \text{if } i \text{ is related to } j \ (i \neq j) \\ \\ 0 & \text{otherwise} \end{cases}$$

- Self-relations are not consider so that the diagonal values  $\boldsymbol{x}_{ii}$  are meaningless

Directed relation: the existence of a tie from i to j does not imply the existence of a tie from j to i (and vice versa)



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### Data



Longitudinal (or panel) network data = M ( $\geq 2$ ) repeated observations on a network

$$x(t_0), x(t_1), \dots, x(t_m), \dots, x(t_{M-1}), x(t_M)$$

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- set of actors  $\mathcal{N} = \{1, 2, \dots, n\}$
- a non reflexive and directed relation  $\ensuremath{\mathcal{R}}$
- actor covariates (gender, age, social status, ...)

## How to model network evolution?



#### Aim

Explain network evolution as a result of

- endogenous variables: structural effects (e.g. reciprocity, transitivity, etc.)
- exogenous variables: actor and dyadic covariates (gender, age, ethnicity, etc.)

simultaneously

Network evolution can be interpreted as the outcome of a **Continuous-time Markov-Chain**...but some assumptions are necessary

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Network evolution can be interpreted as the outcome of a **Continuous-time Markov-Chain**...but some assumptions are necessary

1. *Ties are state*: network ties represent a state with a tendency to endure over time (e.g. friendship, trust, cooperation), rather than a brief event (e.g. telephone calls, e-mail exchanges).

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Network evolution can be interpreted as the outcome of a **Continuous-time Markov-Chain**...but some assumptions are necessary

- Ties are state: network ties represent a state with a tendency to endure over time (e.g. friendship, trust, cooperation), rather than a brief event (e.g. telephone calls, e-mail exchanges).
- 2. Distribution of the process:  $\{X(t), t_0 \le t \le t_M\}$  is a continuous time Markov Chain defined on  $\mathcal{X}$  and  $\mathcal{N}$ .

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State space:  ${\mathfrak X}$  is the set of all possible adjacency matrix (digraphs) defined on the set of actors  ${\mathfrak N}$ 

$$\mathfrak{X} = 2^{n(n-1)} \Rightarrow \mathfrak{X}$$
 is a finite set

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Continuous-time process



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Continuous-time process



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Continuous-time process



Continuous-time process



Latent process: the network evolves in continuous-time but we observed it only at discrete time points

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**Markov property**: the current state of the network determines probabilistically its further evolution

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3. *Opportunity to change*: at a given moment one probabilistically selected actor has the opportunity to change

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4. *Absence of co-occuRrence*: no more than one tie can change at any given moment *t* 

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4. *Absence of co-occurrence*: no more than one tie can change at any given moment *t* 



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- 5. Actor-oriented perspective: the actors control their outgoing ties
  - change in ties are made by the actors who send the ties
  - the actor decide to change one of his outgoing ties according to his position in the network, his attributes and the characteristics of the other actors

#### Aim: maximize a utility function

- actors have complete knowledge about the network and all the other  $\ensuremath{\mathsf{actors}}$
- the maximization is based on immediate returns and not on long-run rewarding (myopic actors)

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# Model definition: assumptions (recap)

- 1. Ties are state
- 2. The evolution process is a continuous-time Markov chain
- 3. At a given moment t one probabilistically selected actor has the opportunity to change

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- 4. No more than one tie can change at any given moment t
- 5. Actor-oriented perspective

#### Consequences of the assumptions

The evolution process can be decomposed into micro-steps: at one randomly determined moment t, one probabilistically selected actor i has the opportunity to change one of his outgoing ties  $x_{ij}$ 

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Micro-step	Continuous-time Markov chain
- the time at which <i>i</i> had the opportunity to change	<ul> <li>the waiting time until the next opportunity for a change made by an actor <i>i</i> (holding time)</li> </ul>

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Micro-step	Continuous-time Markov chain
- the time at which <i>i</i> had the opportunity to change	- the waiting time until the next opportunity for a change made by an actor <i>i</i> (holding time)
- the precise change <i>i</i> made	- the probability of changing the link $x_{ij}$ given the opportunity for changing (jump chain)

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- the precise change <i>i</i> made	- the probability of changing the link x <sub>ij</sub> given the opportunity for changing (jump chain)	

The distribution of the waiting time and the transition matrix of the jump chain are modeled by the *rate function* and the *objective function* respectively.

#### Model definition: rate function

#### How fast is the opportunity for changing?

Waiting time between opportunities of change for actor  $i \sim Exp(\lambda_i)$  $\Rightarrow \lambda_i$  is the expected frequency of changes by actor *i* between observations

Simple specification: all actors have the same rate of change  $\lambda$ 

$$P(i \text{ has the opportunity of change}) = \frac{1}{n} \quad \forall i \in \mathcal{N}$$

#### More complex specification:

- actors may change their ties at different frequencies  $\lambda_i(lpha,x)$
- The parameter of the Exponential distribution is a function of the current state of the network x and the vector of parameter  $\alpha$

$$P(i \text{ has the opportunity of change}) = \frac{\lambda_i(\alpha, x)}{\lambda(\alpha, x)}$$

where

$$\lambda(\alpha, x) = \sum_{i=1}^{n} \lambda_i(\alpha, x)$$

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## Model definition: objective function

#### Which tie is changed?

Change a tie means turning it into its opposite:

$$x_{ij} = 0$$
 is changed into  $x_{ij} = 1$  tie creation

 $x_{ij} = 1$  is changed into  $x_{ij} = 0$  tie deletion

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Given that *i* has the opportunity to change:

Possible choices of $i$	Possible reachable states
n-1 changes	$n-1$ networks $x(i \rightsquigarrow j)$

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Given that *i* has the opportunity to change:

Possible choices of <i>i</i>	Possible reachable states
n-1 changes	$n-1$ networks $x(i \rightsquigarrow j)$
1 non-change	1 network equal to $\times$

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Next state (x')

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To define the non-zero transition probabilities we assume that actors change their ties in order to maximize a utility function.

The Objective function is defined as a linear combination

$$f_i(\beta, x(i \rightsquigarrow j)) = \sum_{k=1}^{K} \beta_k s_{ik}(x(i \rightsquigarrow j)) + U_i(t, x, j)$$

- $s_{ik}(x(i \rightsquigarrow j))$  are effects
- $\beta_k$  are statistical parameters
- $U_i(t,x,j)$  is a random utility term

For a suitable choice of the distribution of  $U_i(t, x, j)$ :

$$p_{ij} = \frac{exp\left(\sum_{k=1}^{K} \beta_k s_{ik}(x(i \rightarrow j))\right)}{\sum_{h=1}^{n} exp\left(\sum_{k=1}^{K} \beta_k s_{ik}(x(i \rightarrow h))\right)}$$

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Observation:  $p_{ii}$  is interpreted as the probability of not changing

Simplified notation:  $x' = x(i \rightsquigarrow j)$ .

**Endogenous effects** = dependent on the network structures

- Outdegree effect

$$s_{i\_out}(x') = \sum_{j} x'_{ij}$$



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$$s_{i\_out}(x') = \sum_{j} x'_{ij}$$



- Reciprocity effect

$$s_{i\_rec}(x') = \sum_{j} x'_{ij} x'_{ji}$$

**Endogenous effects** = dependent on the network structures

- Transitive effect

$$s_{i\_trans}(x') = \sum_{j,h} x'_{ij} x'_{ih} x'_{jh}$$



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**Endogenous effects** = dependent on the network structures

- Transitive effect

$$s_{i\_trans}(x') = \sum_{j,h} x'_{ij} x'_{ih} x'_{jh}$$



- three cycle-effect

$$s_{i\_cyc}(x') = \sum_{j,h} x'_{ij} x'_{jh} x'_{hi}$$



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Example

$$\beta_{out} = -1$$
  $\beta_{rec} = +0.5$   $\beta_{trans} = -0.25$ 

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Example

$$\beta_{out} = -1$$
  $\beta_{rec} = +0.5$   $\beta_{trans} = -0.25$ 



 $p_{11} = 0.146$   $p_{12} = 0.310$   $p_{13} = 0.033$   $p_{14} = 0.511$ 

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#### **Exogenous effects** = related to actor's attributes

#### Example

- Friendship among pupils:

Smoking: non, occasional, regular

Gender: boys, girls

- Trade/Trust (Alliances) among countries:

Geographical area: Europe, Asia, North-America,...

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Worlds: first, Second, Third, Fourth

#### **Exogenous effects** = related to actor's attributes

- covariate-related activity

$$s_{i\_cact}(x) = \sum_{j} x_{ij} z_i$$



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#### **Exogenous effects** = related to actor's attributes

- covariate-related similarity

$$s_{i\_csim}(x) = \sum_{j} x_{ij} \left( 1 - \frac{\left| z_i - z_j \right|}{R_Z} \right)$$

where  $R_Z$  is the range of Z.



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# Which effects must be included in the objective function



Outdegree and Reciprocity must always be included. The choice of the other effects must be determined according to hypotheses derived from theory

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Outdegree and Reciprocity must always be included. The choice of the other effects must be determined according to hypotheses derived from theory

### Example

Friendship: sociological theory suggests that

Theory		Effect
the friend of my friend is also my friend	⇒	transitive effect



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Outdegree and Reciprocity must always be included. The choice of the other effects must be determined according to hypotheses derived from theory

### Example

Friendship: sociological theory suggests that

Theory		Effect
the friend of my friend is also my friend	$\Rightarrow$	transitive effect
girls trust girls boys trust boys	$\Rightarrow$	covariate-related similarity

It is assumed that:

1. the frequencies at which actors have the opportunity to make a change depends on time =  $\lambda$  is not constant over time

*M* time points  $\implies$  we must specify M-1 rate functions

 $\lambda_1, \cdots, \lambda_{M-1}$ 

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*M* time points  $\implies$  we must specify M-1 rate functions

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2. the preferences that drive the choice of the actors have the same intensities over time

 $\beta_1,\cdots,\beta_K$ 

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are constant over time

Consequence: the number of parameters of the SAOM is equal to M - 1 + K



How to interpret the parameter of the SAOM?





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How to interpret the parameter of the SAOM?

- The parameter  $\lambda$  is the expected number of opportunities to change for each actor between two consecutive time points.

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How to interpret the parameter of the SAOM?

- The parameter  $\lambda$  is the expected number of opportunities to change for each actor between two consecutive time points.



- The parameter  $\beta_k$  quantifies the role of each effect in the network evolution.

 $\beta_k=0$  if the corresponding effect plays no role in the network dynamics

 $\beta_k>0$  then there is higher probability of moving into networks where the corresponding effect is higher

 $\beta_k < 0$  there is higher probability of moving into networks where the corresponding effect is lower

### Simulating network evolution

Reproducing a possible series of micro-steps between  $t_0$  and  $t_1$  according to fixed parameter value and the network  $x(t_0)$ .

t =the time

dt = the holding time between consecutive changes

```
Input: x(t_0), \lambda,\beta, n
Output: x^{sim}(t_1)
t \leftarrow 0
x \leftarrow x(t_0)
while condition = TRUE do
    dt \sim Exp(n\lambda)
    i \sim Uniform(1, \ldots, n)
    j \sim Multinomial(p_{i1}, \ldots, p_{in})
    if i \neq j then
    x = x(i \rightarrow j)
    else
    \  \  \, x = x
     t \leftarrow t + dt
x^{sim}(t_1) \leftarrow x
return x^{sim}(t_1)
```

### Simulating network evolution: simulation and unconditional estimation

There are two different stopping rules for the simulations of the network evolution:

1. Unconditional simulation:

the simulations of the network evolution in each time period carry on until a predetermined time length has elapsed (usually until t = 1).

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### Simulating network evolution: simulation and unconditional estimation

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1. Unconditional simulation:

the simulations of the network evolution in each time period carry on until a predetermined time length has elapsed (usually until t = 1).

2. Conditional simulation on the observed number of changes:

simulations run on until the number of different entries between  $x(t_0)$  and the simulated network  $x^{sim}(t_1)$  is equal to the number of entries that differ between  $x(t_0)$  and  $x(t_1)$ 

$$\sum_{\substack{i,j=1\\ i\neq j}}^{n} \left| X_{ij}^{obs}(t_1) - X_{ij}(t_0) \right| = \sum_{\substack{i,j=1\\ i\neq j}}^{n} \left| X_{ij}^{sim}(t_1) - X_{ij}(t_0) \right|$$

This criterion can be generalized conditioning on any other explanatory variable.

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The formulation of the SAOM i depends on M - 1 + K statistical parameters

$$\theta = (\lambda_1, \cdots, \lambda_{M-1}, \beta_1, \cdots, \beta_K)$$

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Aim: estimate  $\theta$ 

Different estimation methods:

- the Method of Moments (MoM)
- the Maximum Likelihood Estimation (MLE)

Definition  $f_X(x;\theta) = \text{probability distribution}$   $\theta \in \Theta \subset \mathbb{R}^p = \text{p-dimensional parameter}$  $X_1, X_2, \dots, X_n = \text{random sample from the probability distribution } f_X(x;\theta)$ 

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The *k*-th *population moment* is:

$$E[X^{k}] = \sum_{x} x^{k} f_{X}(x) \qquad \text{(for the discrete case)}$$
$$E[X^{k}] = \int_{-\infty}^{+\infty} x^{k} f_{X}(x) \qquad \text{(for the continuous case)}$$

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Definition  $f_X(x;\theta) = \text{probability distribution}$   $\theta \in \Theta \subset \mathbb{R}^p = \text{p-dimensional parameter}$  $X_1, X_2, \dots, X_n = \text{random sample from the probability distribution } f_X(x;\theta)$ 

The *k*-th *population moment* is:

$$\begin{split} E[X^k] &= \sum_{x} x^k f_X(x) \qquad (\text{for the discrete case}) \\ E[X^k] &= \int_{-\infty}^{+\infty} x^k f_X(x) \qquad (\text{for the continuous case}) \end{split}$$

The corresponding k-th sample moment is

$$\mu_k = \frac{1}{n} \sum_{i=1}^n X_i^k$$

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To estimate  $\theta$ , one can observe that the theoretical moments of a certain distribution usually depend on the statistical parameters

#### Definition

The method of moment estimators are found by equating the first p population moments to the first p sample moments

 $E[X^{1}] = \mu_{1}$  $E[X^{2}] = \mu_{2}$  $\dots$  $E[X^{p}] = \mu_{p}$ 

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and solving the resulting equations for the unknown parameters.

#### Example

The time to failure of an electronic module used in an automobile engine controller is tested at an elevated temperature to accelerate the failure mechanism.

The time to failure is exponentially distributed with parameter  $\lambda$ .

To estimate the rate parameter  $\lambda$ , eight units are randomly selected and tested, resulting in the following failure time (in hours):

$x_1 = 12.1$	$x_2 = 5.7$	$x_3 = 17.8$	$x_4 = 16.5$
$x_5 = 31.6$	$x_6 = 7.7$	$x_7 = 11.9$	$x_8 = 22.7$

What is the estimate for  $\lambda$  according to the observed data and the the MoM?

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### Example

The first population moment of the Exponential random variable is

$$E[X] = \frac{1}{\lambda}$$

and the corresponding sample moment is

$$\mu_1 = \frac{1}{n} \sum_{i=1}^n X_i$$

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and the corresponding sample moment is

$$\mu_1 = \frac{1}{n} \sum_{i=1}^n X_i$$

According to the MoM, the estimator for the parameter  $\lambda$  is:

$$\frac{1}{\lambda} = \frac{1}{n} \sum_{i=1}^{n} X_i \qquad \Leftrightarrow \qquad \lambda = \frac{n}{\sum_{i=1}^{n} X_i}$$

and the corresponding estimate is

$$\widehat{\lambda} = \frac{n}{\sum\limits_{i=1}^{n} x_i} = \frac{8}{126} = 0.063$$

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The principle of the MoM can be easily generalized to any set of p functions  $s_k(X)$ , k = 1, ..., p

Population moment

 $E[s_k(X)] = \sum_{x} s_k(x) f_X(x) \quad \text{(for the discrete case)}$  $E[s_k(X)] = \int_{-\infty}^{+\infty} s_k(x) f_X(x) \quad \text{(for the continuous case)}$ 

Sample moment

$$\gamma_k = \frac{1}{n} \sum_{i=1}^n s_k(X_i)$$

 $s_k(X)$  are called statistics

They **must be sensitive** to the parameter  $\theta$ , i.e. higher values of  $\theta$  lead to higher values of s(X).

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#### Remark

- An *estimator* is a function of the sample, e.g.  $\frac{n}{\sum_{i=1}^{n} X_i}$
- An estimate is the realized value of an estimator, e.g.  $\frac{n}{\sum\limits_{i=1}^{n} x_i}$
- The estimate of a parameter varies according to the selected sample. Thus, we usually associate to an estimator its standard error.

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#### Example

We assume to randomly select and test other eight electronic modules, resulting in the following failure time (in hours):

$x_1 = 9.5$	$x_2 = 7.2$	$x_3 = 13.4$	$x_4 = 10.2$
$x_5 = 15.0$	$x_6 = 16.3$	$x_7 = 13.9$	<i>x</i> <sub>8</sub> = 34.5

The new estimate for the parameter  $\lambda$  is

$$\widehat{\lambda} = \frac{n}{\sum\limits_{i=1}^{n} x_i} = \frac{8}{120} = 0.067$$

This value is close to the previous but it is not the same!

Aim: estimate  $\theta$  using the MoM

$$\theta = (\lambda_1, \cdots, \lambda_{M-1}, \beta_1, \cdots, \beta_K)$$

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In practice:

- find M 1 + K statistics
- set the theoretical expected value of each statistic equal to its sample counterpart

- solve the resulting system of equations with respect to  $\theta$ .

Aim: estimate  $\theta$  using the MoM

$$\theta = (\lambda_1, \cdots, \lambda_{M-1}, \beta_1, \cdots, \beta_K)$$

In practice:

- find M 1 + K statistics
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- solve the resulting system of equations with respect to  $\theta$ .

Which statistics are suitable for estimating  $\theta$ ?

Aim: estimate  $\theta$  using the MoM

$$\theta = (\lambda_1, \cdots, \lambda_{M-1}, \beta_1, \cdots, \beta_K)$$

In practice:

- find M 1 + K statistics
- set the theoretical expected value of each statistic equal to its sample counterpart
- solve the resulting system of equations with respect to  $\theta$ .



For simplicity, let us assume to have observed a network at two time points  $t_0$  and  $t_1$  and to condition the estimation on the first observation  $x(t_0)$ 

- The rate parameter  $\lambda$  describes the frequencies at which changes happen.

$$s_{\lambda}(X(t_1), X(t_0)|X(t_0) = x(t_0)) = \sum_{\substack{i,j=1 \ i \neq j}}^n |X_{ij}(t_1) - X_{ij}(t_0)|$$

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 $\Rightarrow$  a high value of  $\lambda$  leads to a high value of  $s_\lambda(X(t_1),X(t_0)|X(t_0)=x(t_0))$ 

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- The parameter  $\beta_k$  quantifies the role played by each effect in the network evolution.

$$s_k(X(t_1)|X(t_0) = x(t_0)) = \sum_{i=1}^n s_{ik}(X(t_1))$$





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$$s_k(X(t_1)|X(t_0) = x(t_0)) = \sum_{i=1}^n s_{ik}(X(t_1))$$

E.g.: let us consider the parameter  $\beta_{out}$ . The corresponding statistic is

$$s_{out}(X(t_1)|X(t_0) = x(t_0)) = \sum_{i=1}^n s_{i\_out}(X(t_1)) = \sum_{i=1}^n \sum_{j=1}^n x_{ij}(t_1)$$

	$\beta_{out} = -2.5$	$\beta_{out} = -2$	$\beta_{out} = -1.5$
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 $\Rightarrow$  a high value of  $\beta_{out}$  leads to a high value of  $s_{out}(X(t_1)|X(t_0)=x(t_0))$ 

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Generalizing to M-1 periods:

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- Statistics for the rate function parameters

$$s_{\lambda_1}(X(t_1), X(t_0)|X(t_0) = x(t_0)) = \sum_{\substack{i,j=1 \ i \neq j}}^n |X_{ij}(t_1) - X_{ij}(t_0)|$$

 $s_{\lambda_{M-1}}(X(t_M), X(t_{M-1})|X(t_{M-1}) = x(t_{M-1})) = \sum_{\substack{i,j=1\\ i \neq j}}^{n} |X_{ij}(t_M) - X_{ij}(t_{M-1})|$ 

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Generalizing to M-1 periods:

. . .

- Statistics for the rate function parameters

$$s_{\lambda_1}(X(t_1), X(t_0) | X(t_0) = x(t_0)) = \sum_{\substack{i,j=1\\ i \neq j}}^n |X_{ij}(t_1) - X_{ij}(t_0)|$$

$$s_{\lambda_{M-1}}(X(t_M), X(t_{M-1})|X(t_{M-1}) = x(t_{M-1})) = \sum_{\substack{i,j=1\\i \neq j}}^{n} |X_{ij}(t_M) - X_{ij}(t_{M-1})|$$

- Statistics for the objective function parameters:

$$\sum_{m=1}^{M-1} s_{mk}(X(t_m)|X(t_{m-1}) = x(t_{m-1})) = \sum_{m=1}^{M-1} s_{mk}(X(t_m), X(t_{m-1}))$$

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The MoM estimator for  $\theta$  is defined as the solution of the system of M + K - 1 equations

$$\begin{cases} E_{\theta} \left[ s_{\lambda_m}(X(t_m), X(t_{m+1}) | X(t_m) = x(t_m)) \right] = s_{\lambda_m}(x(t_1), x(t_0)) \\ E_{\theta} \left[ \sum_{m=1}^{M-1} s_{mk}(X(t_{m+1}) | X(t_m) = x(t_m)) \right] = \sum_{m=1}^{M-1} s_{mk}(x(t_{m+1}), x(t_m)) \end{cases}$$

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with  $m = 1, \dots, M-1$  and  $k = 1, \cdots, K$ 

### Example

Let us assume to have observed a network at M = 3 time points



and to model network evolution considering the outdegree and the reciprocity effects.

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$$\theta = (\lambda_1, \lambda_2, \beta_{out}, \beta_{rec})$$

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## Example

Statistics:

$$s_{\lambda_1}(X(t_1), X(t_0)|X(t_0) = x(t_0)) = \sum_{\substack{i,j=1 \ i \neq j}}^4 |X_{ij}(t_1) - X_{ij}(t_0)|$$

$$s_{\lambda_2}(X(t_2), X(t_1)|X(t_1) = x(t_1)) = \sum_{\substack{i,j=1 \ i \neq j}}^4 |X_{ij}(t_2) - X_{ij}(t_1)|$$

$$\sum_{m=1}^{2} s_{m-out}(X(t_m)|X(t_{m-1}) = x(t_{m-1})) = \sum_{m=1}^{2} \sum_{i=1}^{4} \sum_{j=1}^{4} X_{ij}(t_m)$$

$$\sum_{m=1}^{2} s_{m\_rec} \left( X(t_m) | X(t_{m-1}) = x(t_{m-1}) \right) = \sum_{m=1}^{2} \sum_{i=1}^{4} \sum_{j=1}^{4} X_{ij}(t_m) X_{ji}(t_m)$$

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Example



Observed values of the statistics:

 $s_{\lambda_1} = 5$   $s_{\lambda_2} = 4$  $\sum_{m=1}^{2} s_{m\_out} = 6 + 8 = 14$   $\sum_{m=1}^{2} s_{m\_rec} = 2 + 3 = 5$ 

#### Example

We look for the value of  $\boldsymbol{\theta}$  that satisfies the system:

$$\begin{cases} E_{\theta} \left[ s_{\lambda_{1}}(X(t_{1}), X(t_{0}) | X(t_{0}) = x(t_{0})) \right] = 5 \\ E_{\theta} \left[ s_{\lambda_{2}}(X(t_{2}), X(t_{1}) | X(t_{1}) = x(t_{1})) \right] = 4 \\ E_{\theta} \left[ \sum_{m=1}^{2} s_{m\_out} \left( X(t_{m}) | X(t_{m-1}) = x(t_{m-1}) \right) \right] = 14 \\ E_{\theta} \left[ \sum_{m=1}^{2} s_{m\_rec} \left( X(t_{m}) | X(t_{m-1}) = x(t_{m-1}) \right) \right] = 5 \end{cases}$$

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Simplified notation:

- S: (M 1 + K)-dimensional vector of statistics
- s: (M-1+K)-dimensional vector of the observed values of the statistics

Consequently, the system of moment equations can be written as

$$E_{\theta}[S] = s$$

or equivalently as

 $E_{\theta}[S-s]=0$ 

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Problem: analytical and usual numerical procedures cannot be applied to solve this system, but we can use a stochastic approximation method

#### Definition

*Stochastic approximation methods* are a family of iterative stochastic optimization algorithms that attempt to find zeros or extrema of functions which cannot be computed in an analytical way.

- 1. The expected values of the statistics are approximated via Monte Carlo methods  $\Rightarrow$  stochastic
- 2. The value of  $\theta$  is iteratively updated according to the "distance" between the approximated expected values and the observed values.

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#### 1. Approximation of the expected value $E_{\theta}[S]$

1. Given  $x(t_0)$  and  $\theta$ , simulate the sequence of the observed networks at time  $t_1, \ldots, t_M q$  times. Denote these sequences by

$$x^{(1)}(t_1), x^{(1)}(t_2), \ldots, x^{(1)}(t_M)$$

$$x^{(q)}(t_1), x^{(q)}(t_2), \ldots, x^{(q)}(t_M)$$

- 2. For each sequence compute the value  $S^{(l)}$  assumed by S
- 3. Approximate the expected value by

$$\overline{S} = \frac{1}{q} \sum_{l=1}^{q} S^{(l)}$$

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3. Approximate the expected value by

$$\overline{S}_{out} = \frac{1}{q} \sum_{i=1}^{q} S_{out}^{(i)} \to E_{\theta}[S_{out}]$$
#### 2. Updating the value of $\theta$

The value of  $\theta$  is updated by the Robbins-Monro step:

$$\widehat{\theta}_{i+1} = \widehat{\theta}_i - a_i D^{-1} (\overline{S}_i - s)$$

where

-  $a_i$  is a sequence of positive numbers such that:

$$\lim_{(i\to\infty)}a_i=0\qquad \sum_{i=1}^\infty a_i=\infty\qquad \sum_{i=1}^\infty a_i^2<\infty$$

- *D* denotes the diagonal matrix of the first order derivative matrix of *S* with respect to  $\theta$ :

$$D = \frac{\partial}{\partial \theta} E_{\theta}[S|X(t_0) = x(t_0)]$$

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2. Updating the value of  $\theta$ 

### Example

We look for the value of  $\theta$  that satisfies the system:

$$\begin{cases} E_{\theta} \left[ s_{\lambda_{1}}(X(t_{1}), X(t_{0}) | X(t_{0}) = x(t_{0})) \right] = 5 \\ E_{\theta} \left[ s_{\lambda_{2}}(X(t_{2}), X(t_{1}) | X(t_{1}) = x(t_{1})) \right] = 4 \\ E_{\theta} \left[ \sum_{m=1}^{2} s_{m\_out} \left( X(t_{m}) | X(t_{m-1}) = x(t_{m-1}) \right) \right] = 14 \\ E_{\theta} \left[ \sum_{m=1}^{2} s_{m\_rec} \left( X(t_{m}) | X(t_{m-1}) = x(t_{m-1}) \right) \right] = 5 \end{cases}$$

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2. Updating the value of  $\boldsymbol{\theta}$ 

## Example

- Initial guess  $\theta_0 = (4.5, 3.2, -0.2, 0.9)$
- Simulate the network evolution 1000 times according to  $\theta_0$

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- Approximation of the expected values  $E_{ heta_0}[\overline{S}]$ 

 $\overline{S}_{\lambda_1} = 6.211 \qquad \overline{S}_{\lambda_2} = 4.567 \\ \overline{S}_{\beta_{out}} = 13.806 \qquad \overline{S}_{\beta_{rec}} = 4.702$ 

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### - Approximation of the moment equation

$$\begin{split} \overline{S}_{\lambda_1} - 5 &= 1.211 & \overline{S}_{\lambda_2} - 4 = 0.567 \\ \overline{S}_{\beta_{out}} - 14 &= -0.194 & \overline{S}_{\beta_{rec}} - 5 = -0.298 \end{split}$$

2. Updating the value of  $\theta$ 

## Example

- Initial guess  $\theta_0 = (4.5, 3.2, -0.2, 0.9)$
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- Approximation of the moment equation
  - $\overline{S}_{\lambda_1} 5 = 1.211 \qquad \overline{S}_{\lambda_2} 4 = 0.567 \\ \overline{S}_{\beta_{out}} 14 = -0.194 \qquad \overline{S}_{\beta_{rec}} 5 = -0.298$

The Robbins-Monro step

$$\widehat{\theta}_{i+1} = \widehat{\theta}_i - a_i D^{-1} (\overline{S}_i - s)$$

suggests how to modify the parameter value to satisfy the moment equation through the difference  $(\overline{S}_i - s)$ 

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2. Updating the value of  $\theta$ 

## Example

- Guess  $\theta_1 = (4.1, 2.9, -0.2, 0.9)$
- Simulate the network evolution 1000 times according to  $\theta_1$

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- Approximation of the expected values  $E_{ heta_1}[\overline{S}]$ 

$$\overline{S}_{\lambda_1} = 5.345 \qquad \overline{S}_{\lambda_2} = 4.215 \\ \overline{S}_{\beta_{out}} = 13.813 \qquad \overline{S}_{\beta_{rec}} = 4.759$$

2. Updating the value of  $\theta$ 

## Example

- Guess  $\theta_1 = (4.1, 2.9, -0.2, 0.9)$
- Simulate the network evolution 1000 times according to  $\theta_1$

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- Approximation of the moment equation
  - $\overline{S}_{\lambda_1} 5 = 0.345 \qquad \overline{S}_{\lambda_2} 4 = 0.215 \\ \overline{S}_{\beta_{out}} 14 = -0.187 \qquad \overline{S}_{\beta_{rec}} 5 = -0.241$

2. Updating the value of  $\theta$ 

## Example

- Guess  $\theta_i = (3.87, 2.56, -0.11, 0.87)$
- Simulate the network evolution 1000 times according to  $\theta_1$

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- Approximation of the expected values  $E_{\theta_i}[\overline{S}]$ 

$$\overline{S}_{\lambda_1} = 5.055 \qquad \overline{S}_{\lambda_2} = 4.018 \\ \overline{S}_{\beta_{out}} = 14.012 \qquad \overline{S}_{\beta_{rec}} = 4.974$$

2. Updating the value of  $\boldsymbol{\theta}$ 

## Example

- Guess  $\theta_i = (3.87, 2.56, -0.11, 0.87)$
- Simulate the network evolution 1000 times according to  $\theta_1$

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**Phase 1**: given the network at time  $t_0$  and an initial guess  $\theta_0$  for  $\theta$  a small number  $q_1$  of steps are made to estimate D.

1. Network evolution is simulated from  $\theta_0$  and the values  $S_{i0}$  are computed

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- 2. Network evolution is simulated from  $\theta_0 + \epsilon_j e_j$  and the values  $S_{ij}$  are computed where
  - e<sub>j</sub> is the j-th unit vector
  - 0.1 <  $\epsilon_j < 1$

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- 5. Estimate  $E_{\theta}[S]$  and D by the Monte Carlo method

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  $\widehat{d}_j = rac{1}{q_1} \sum_{i=1}^{q_1} d_j$ 

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6. A new value of  $\theta$  is estimated via the Robbins-Monro step with  $a_i = 1$ :

$$\widehat{\theta}_{q_1} = \theta_0 - \widehat{D}^{-1}(\overline{S} - s)$$

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**Phase 2**: the main phase, consisting of c sub-phases. Each sub-phase h iterate the Robbins-Monro step for at most  $q_h$  step.

1. Generate the network evolution according to  $\hat{\theta}_i$  and compute  $S_{ih}$ 

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$$\widehat{\theta}_{i+1} = \widehat{\theta}_i - a_h \widehat{D}^{-1} (\overline{S}_i - s)$$

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3. Repeat steps 1. and 2. until  $i > q_h$  or  $(S_{ih} - s)(S_{(i-1)h} - s) < 0$ 

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- 4. Compute

$$\widehat{\theta}_h = \frac{1}{i} \sum_i \widehat{\theta}_i$$

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- 5.  $a_{h+1} = \frac{a_h}{2}$  is the eventual estimate for  $\theta$
- $\widehat{\theta_c}$  is the eventual estimate for  $\theta$

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**Phase 3**: a number of  $q_3$  simulations is used to evaluate the convergence of the algorithm and the accuracy of the estimator.

- Phase 3 requires the computation of D, thus it is similar to Phase 1.

## The Maximum-likelihood estimation (MLE)

The model assumptions allow to decompose the process in a series of micro-steps:

$$\{(T_r, i_r, j_r), r = 1, \ldots, R\}$$

where

- $T_r$  is the time point for an opportunity for change
- *i<sub>r</sub>* denotes the actor who has the opportunity to change
- $j_r$  is the actor towards whom the tie is changed

We denote by *R* the total number of micro-steps between  $t_0$  and  $t_1$  and we assume that the time point  $T_r$  are ordered increasingly:

$$t_0 = T_0 < T_1 < \ldots < T_R < t_1$$

## The Maximum-likelihood estimation (MLE)

#### Definition

Given the sequence of micro-steps  $\{(T_r, i_r, j_r), r = 0, ..., R\}$ , the likelihood function of the network evolution process is defined by:

$$L(\theta) = \prod_{r=1}^{R} P((T_r, i_r, j_r))$$

Then, the estimate for  $\theta$  is the vector of values  $\widehat{\theta}$  such that:

$$\widehat{\theta} = \arg \max_{\theta \in \Theta} L(\theta)$$

or equivalently, the vector of values  $\widehat{\theta}$  such that:

$$U(\theta) = \frac{\partial}{\partial \theta} \log(L(\theta)) = 0$$

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where  $U(\theta)$  is the score function.

Problem: we cannot observe the complete data and the likelihood of the observed data  $(x(t_1), \ldots, x(t_M))$  conditional on  $x(t_0)$ .

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the parameter estimation requires finding the root of a system of equations in which the functions cannot be computed analytically  $\Rightarrow$  a stochastic approximation method must be applied.

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The idea is to augment the observed data so that an easily computable likelihood is obtained.

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Note: the data augmentation can be done separately for each time period  $(t_{m-1}, t_m)$ . It is not restrictive to describe it only for two observations  $x(t_0)$  and  $x(t_1)$ .

### Definition

The *augmented data* (or *sample path*) consist of *R* and of the sequence of tie changes that brings the network from  $x(t_0)$  to  $x(t_1)$ 

$$(i_1,j_1),\ldots,(i_R,j_R)$$

Formally:

$$\underline{v} = \{(i_1, j_1), \ldots, (i_R, j_R)\} \in \mathcal{V}$$

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We can approximate the likelihood function (and then the score function) of the observed data using the probability of  $\underline{v}$ 

$$P(\underline{v}|x(t_0),x(t_1)) \propto \frac{(n\lambda)^R}{R!} e^{-n\lambda} \prod_{r=1}^R \frac{1}{\lambda} p_{i_r j_r}(\beta,x(T_{r-1}))$$

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The augmented data are sampled through a Markov chain simulation defined on  $\ensuremath{\mathcal{V}}$ 

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### Definition

The *Metropolis-Hastings algorithm* is defined by the following transition probabilities:

1. Given  $\underline{v}_i = \underline{v}$ , generate  $\underline{\widetilde{v}}$  form the proposal distribution  $u(\underline{\widetilde{v}} | \underline{v}_i)$ 

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### Definition

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- 2. Take

$$\underline{\nu}_{i+1} = \left\{ \begin{array}{ll} \underline{\widetilde{\nu}} & \text{with probability} \quad \rho(\underline{\widetilde{\nu}},\underline{\nu}) \\ \\ \underline{\nu} & \text{with probability} \quad 1 - \rho(\underline{\widetilde{\nu}},\underline{\nu}) \end{array} \right.$$

where

$$\rho(\underline{\widetilde{v}},\underline{v}) = \min\left\{\frac{P(\underline{\widetilde{v}})u(\underline{v}\,|\underline{\widetilde{v}}\,)}{P(\underline{v})u(\underline{\widetilde{v}}\,|\underline{v}\,)}, 1\right\}$$

The transition probabilities of the chain generate by the Metropolis-Hastings algorithm are given by  $\rho(\underline{\widetilde{v}},\underline{v})u(\underline{\widetilde{v}}\,|\underline{v}\,)$ 

The proposal distribution  $u(\underline{\tilde{v}}|\underline{v})$  assigns non-null probabilities to the following changes:

1. Pairwise deletions: one pair of indices  $r_1$  and  $r_2$  such that  $(i_{r_1}, j_{r_1}) = (i_{r_2}, j_{r_2})$  is selected and the corresponding pairs  $(i_{r_1}, j_{r_1})$  and  $(i_{r_2}, j_{r_2})$  are deleted from  $\underline{v}$ 

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Example



 $\underline{v} = (2,4) (2,3) (1,1) (4,2) (3,2) (1,4) (2,4) (2,3) (1,3) (2,4) (3,3)$ 

- Select at random  $(r_1, r_2)$  in  $\{(1,7)(1,10)(2,8)\}$ , e.g.  $(r_1, r_2) = (1,7)$
- Delete the elements (2,4)

 $\underline{\widetilde{v}} = (2,3) (1,1) (4,2) (3,2) (1,4) (2,3) (1,3) (2,4) (3,3)$ 

The proposal distribution  $u(\underline{\tilde{v}}|\underline{v})$  assigns non-null probabilities to the following changes:

2. Pairwise insertions: one pair of  $(i,j) \in \mathbb{N}^2$  and two indices  $r_1$  and  $r_2$  are randomly chosen. The element (i,j) is inserted in  $\underline{v}$  immediately before  $r_1$  and  $r_2$ .

The proposal distribution  $u(\underline{\tilde{\nu}} | \underline{\nu})$  assigns non-null probabilities to the following changes:

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 $\underline{v} = (2,4) (2,3) (1,1) (4,2) (3,2) (1,4) (2,4) (2,3) (1,3) (2,4) (3,3)$ 

- Select at random (i, j) and  $(r_1, r_2)$ , e.g.  $i = 4, j = 1, r_1 = 5, r_2 = 7$
- Insert the elements (4,1) before  $r_1 = 5$  and  $r_2 = 7$

 $\widetilde{\underline{v}} = (2,4) (2,3) (1,1) (4,2) (4,1) (3,2) (1,4) (4,1) (2,4) (2,3) (1,3) (2,4) (3,3)$
The proposal distribution  $u(\underline{\tilde{v}}|\underline{v})$  assigns non-null probabilities to the following changes:

3. Single deletion: one pair  $(i_r, j_r)$  satisfying  $i_r = j_r$  is randomly selected and deleted from  $\underline{v}$ 

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#### Example



 $\underline{v} = (2,4) (2,3) (1,1) (4,2) (3,2) (1,4) (2,4) (2,3) (1,3) (2,4) (3,3)$ 

- Select at random r in  $\{3,11\}$ , e.g. r = 11
- Delet the elements (3,3)

 $\widetilde{\underline{v}} = (2,4) \ (2,3) \ (1,1) \ (4,2) \ (3,2) \ (1,4) \ (2,4) \ (2,3) \ (1,3) \ (2,4)$ 

The proposal distribution  $u(\underline{\tilde{v}}|\underline{v})$  assigns non-null probabilities to the following changes:

4. Single insertion: one actor  $i \in \mathbb{N}$  and an index r are selected. The element (i, i) is inserted immediately before r

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Example



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- Select at random  $i \in \mathcal{N}$  and r, e.g. i = 4 r = 6
- Insert the elements (4,4) before r = 6

 $\underline{\tilde{v}} = (2,4) (2,3) (1,1) (4,2) (3,2) (1,4) (4,4) (2,4) (2,3) (1,3) (2,4) (3,3)$ 

The proposal distribution  $u(\underline{\tilde{v}}|\underline{v})$  assigns non-null probabilities to the following changes:

5. *Permutations*: for randomly chosen indices  $r_1 < r_2$ , the sequence  $(i_{r_1}, j_{r_1}), \ldots, ((i_{r_2}, j_{r_2}))$  is randomly permuted

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- Select at random  $(r_1, r_2)$  and r, e.g.  $r_1 = 2, r_2 = 5$
- Permute the sequence  $(i_2, j_2), \ldots, (i_5, j_5)$

 $\underline{v} = (2,4)$  (1,1) (2,3) (3,2) (4,2) (1,4) (4,4) (2,4) (2,3) (1,3) (2,4) (3,3)

#### Theorem

The Metropolis-Hastings algorithm leads to an irreducible, aperiodic and reversible Markov-chain.

#### Proof

- The Markov chain is irreducible. Pairwise deletions and insertions and single deletion and insertion are sufficient for all  $\underline{v} \in$  to communicate.
- The Markov chain is aperiodic.

The graph associated to the resulting Markov-chain contains all the loops and thus the greatest common divisor of all cycles is one.

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- The Markov chain is reversible. The detailed balance condition:

$$\rho(\underline{\widetilde{v}},\underline{v})u(\underline{\widetilde{v}}|\underline{v})P(\underline{v}) = \rho(\underline{v},\underline{\widetilde{v}})u(\underline{v}|\underline{\widetilde{v}})P(\underline{\widetilde{v}})$$

is satisfied.

$$\rho(\widetilde{\underline{v}}, \underline{v})u(\widetilde{\underline{v}} | \underline{v})P(\underline{v}) = \min\left\{\frac{P(\widetilde{\underline{v}})u(\underline{v} | \widetilde{\underline{v}})}{P(\underline{v})u(\widetilde{\underline{v}} | \underline{v})}, 1\right\}u(\widetilde{\underline{v}} | \underline{v})P(\underline{v}) =$$

$$= \min\left\{\frac{P(\widetilde{\underline{v}})u(\underline{v} | \widetilde{\underline{v}})}{u(\widetilde{\underline{v}} | \underline{v})}, P(\underline{v})\right\}u(\widetilde{\underline{v}} | \underline{v}) =$$

$$= \min\left\{\frac{u(\underline{v} | \widetilde{\underline{v}})}{u(\widetilde{\underline{v}} | \underline{v})}, \frac{P(\underline{v})}{P(\widetilde{\underline{v}})}\right\}u(\underline{\widetilde{v}} | \underline{v})P(\widetilde{\underline{v}}) =$$

$$= \min\left\{1, \frac{P(\underline{v})u(\widetilde{\underline{v}} | \underline{v})}{P(\widetilde{\underline{v}})u(\underline{v} | \widetilde{\underline{v}})}\right\}u(\underline{v} | \widetilde{\underline{v}})P(\widetilde{\underline{v}}) =$$

$$= \rho(\underline{v}, \widetilde{\underline{v}})u(\underline{v} | \underline{\widetilde{v}})P(\widetilde{\underline{v}})$$

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The ML estimation algorithm can be sketched in the following way:

- 1. For each m = 1, ..., M-1 makes a large number of Metropolis-Hastings steps yielding  $v^{(i)} = (v_1^{(i)}, ..., v_{M-1}^{(i)})$
- 2. Compute the score function:

$$\frac{\partial}{\partial \theta} log(L(\widehat{\theta}_i; x; v_m^{(i)}))$$

3. Update the parameter estimate using the Robbins-Monro step

$$\theta_{i+1} = \theta_i + a_i D^{-1} U(L(\widehat{\theta}_i; x; v_m^{(i)}))$$

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The estimate  $\widehat{\theta}$  is calculated as the average of the  $\theta_{i+1}$  values generated by this algorithm.

### Parameter estimation

The Robbins-Monro algorithm and the ML estimation are implemented in the R library *RSiena* (Simulation Investigation for Empirical Network Analysis)

You need to load the following libraries:

- 1. library(snow)
- 2. library(rlecuyer)
- 3. library(RSiena)

The R script "estimation.R" contains the R commands to implement the estimation procedure in R and the folder "tfls.zip" includes the data files.

Example data: an excerpt from the "Teenage Friends and Lifestyle Study" data set:

- Networks: relation = friendship

actors = 129 pupils present at all three measurement points

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	Estimates	s.e.	t-score
Rate parameters:			
Rate parameter period 1	8.5948	(0.7091)	
Rate parameter period 2	7.2115	(0.5751)	
Other parameters:			
outdegree (density)	-2.4147	(0.0387)	-62.3875
reciprocity	2.7106	(0.0811)	33.4061

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Rate parameter: expected frequency, between two consecutive network observations, with which actors get the opportunity to change a network tie

- about 9 opportunities for change in the first period
- about 7 opportunities for change in the second period

The estimated rate parameters will be higher than the observed number of changes per actor (why?)

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#### Interpreting the objective function parameters:

The parameter  $\beta_k$  quantifies the role of the effect  $s_k$  in the network evolution.

 $\beta_k = 0 \ s_k$  plays no role in the network dynamics

 $\beta_k > 0$  higher probability of moving into networks where  $s_k$  is higher

 $\beta_k < 0$  higher probability of moving into networks where  $\mathbf{s}_k$  is lower



Which  $\beta_k$  are "significantly" different from 0? E.g.  $\beta_{rec} = 0.13$  is "significantly" different from 0?

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Hypothesis test:

- 1. State the hypotheses.
  - The *null hypothesis*  $(H_0)$  states that the observed increase or decrease in the number of network configurations related to a certain effect results purely from chance.

$$H_0: \beta_k = 0$$

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$$H_0:\beta_k=0$$

- The *alternative hypothesis*  $(H_1)$  states that the observed increase or decrease in the number of network configurations related to a certain effect is influenced by some non-random cause.

$$H_1: \beta_k \neq 0$$

Hypothesis test:

2. Define a decision rule

$$\begin{cases} \left| \frac{\beta_k}{s.e.(\beta_k)} \right| \ge 2 & \text{reject } H_0 \\ \left| \frac{\beta_k}{s.e.(\beta_k)} \right| < 2 & \text{fail to reject } H_0 \end{cases}$$

The logic behind this decision rule is based on the standard error concept.

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## Example

Is the value  $\beta_{rec} = 0.13$  far enough from 0?

If  $s.e.(\beta_{\it rec})=0.9,$  a more or less plausible set of values that the parameter can assume is approximately

[0.04, 0.22]

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$$\left|\frac{\beta_{\textit{rec}}}{s.e.(\beta_{\textit{rec}})}\right| = \left|\frac{0.13}{0.9}\right| = 0.14 < 2$$

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 $\beta_{rec}$  is not significantly different from 0

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Objective function parameters:

- outdegree parameter: the observed networks have low density

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Objective function parameters:

- outdegree parameter: the observed networks have low density
- reciprocity parameter: strong tendency towards reciprocated ties

In more detail

$$\beta_{out} \sum_{j=1}^{n} x_{ij} + \beta_{rec} \sum_{j=1}^{n} x_{ij} x_{ji} = -2.4147 \sum_{j=1}^{n} x_{ij} + 2.7106 \sum_{j=1}^{n} x_{ij} x_{ji}$$

Adding a reciprocated tie (i.e., for which  $x_{ji} = 1$ ) gives

-2.4147 + 2.7106 = 0.2959

while adding a non-reciprocated tie (i.e., for which  $x_{ji} = 0$ ) gives

-2.4147

Conclusion: reciprocated ties are valued positively and non-reciprocated ties are valued negatively by actors

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#### Specifying the objective function

In friendship context, sociological theory suggests that:

- friendship relations tend to be reciprocated  $\rightarrow$  reciprocity effect



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- the statement "the friend of my friend is also my friend" is almost always true

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#### Specifying the objective function

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- friendship relations tend to be reciprocated  $\rightarrow$  reciprocity effect



- the statement "the friend of my friend is also my friend" is almost always true  $\rightarrow$  transitive triplets effect



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Specifying the objective function

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This effect must be controlled for the sender and receiver effects of the covariate.

- Covariate ego effect



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	Estimates	s.e.	t-score
Rate parameters:			
Rate parameter period 1	10.6809	(1.0425)	
Rate parameter period 2	9.0116	(0.8386)	
Other parameters:			
outdegree (density)	-2.8597	(0.0608)	-47.0288
reciprocity	1.9855	(0.0876)	22.6765
transitive triplets	0.4480	(0.0257)	17.4558
sex alter	-0.1513	(0.0980)	-1.5445
sex ego	0.1571	(0.1072)	1.4659
sex similarity	0.9191	(0.1076)	8.5440
smoke alter	0.1055	(0.0577)	1.8272
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- outdegree parameter: the observed networks have low density

- reciprocity parameter: strong tendency towards reciprocated ties
- transitivity parameter: preference for being friends with friends' friends

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- sex alter: gender does not affect actor popularity

- sex ego: gender does not affect actor activity
- sex similarity: tendency to choose friends with the same gender

- Gender: coded with 1 for boys and with 2 for girls.

- Gender: coded with 1 for boys and with 2 for girls.
- All actor covariates are centered:  $\overline{z} = 1.434$  is the mean of the covariate

$$z_i - \overline{z} = \begin{cases} -0.434 & \text{ for boys} \\ \\ 0.566 & \text{ for girls} \end{cases}$$

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- The contribution of x<sub>ii</sub> to the objective function is

$$\beta_{ego}(z_i - \overline{z}) + \beta_{alter}(z_j - \overline{z}) + \beta_{same} \mathbb{I}\{z_i = z_j\} =$$

$$= 0.1571(z_i - \overline{z}) - 0.1513(z_j - \overline{z}) + 0.9191\mathbb{I}\{z_i = z_j\}$$

where  $\mathbb{I}\{z_i = z_j\}$  is the indicator function

$$\mathbb{I}\{z_i = z_j\} \begin{cases} 1 & z_i = z_j \\ 0 & \text{otherwise} \end{cases}$$

	Male	Female
Male	0.9166	0.1546
Female	-0.1538	0.9224

Table: Gender-related contributions to the objective function

#### Conclusions:

- preference for similar alters
- the negative value associated to the the single tie form a girl to a boy, suggests that girls seem not to like male friends.

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- smoke alter: smoking behavior does not affect actor popularity
- smoke ego: smoking behavior not affect actor activity
- smoke similarity: tendency to choose friends with the same smoking behavior

- Smoking behavior: coded with 1 for "no", 2 for "occasional", and 3 for "regular" smokers.

- Smoking behavior: coded with 1 for "no", 2 for "occasional", and 3 for "regular" smokers.
- The smoking covariate is centered:  $\overline{z} = 1.310$  is the mean of the covariate

$$z_i - \overline{z} = \begin{cases} -0.310 & \text{ for no smokers} \\ 0.690 & \text{ for occasional smokers} \\ 1.690 & \text{ for regular smokers} \end{cases}$$

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- The contribution of  $x_{ij}$  to the objective function is

$$\beta_{ego}(z_i - \overline{z}) + \beta_{alter}(z_j - \overline{z}) + \beta_{same} \left( 1 - \frac{|z_i - z_j|}{R_z} - sim_z \right) = 0.0714(z_i - \overline{z}) + 0.1055(z_j - \overline{z}) + 0.3724 \left( 1 - \frac{|z_i - z_j|}{2} - 0.7415 \right)$$

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where  $R_z = z_{max} - z_{min}$ 

	no	occasional	regular
no	0.0414	-0.0734	-0.1882
occasional	-0.0393	0.2183	0.1035
regular	-0.1200	0.1376	0.3952

Table: Smoking-related contributions to the objective function

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### Conclusions:

- preference for similar alters
- this tendency is strongest for high values on smoking behavior

# Outline

Introduction

The Stochastic actor-oriented model

Extending the model: analyzing the co-evolution of networks and behavior

Motivation Selection and influence Model definition and specification Simulating the co-evolution of networks and behavior Parameter estimation Parameter interpretation

# Networks are dynamic by nature: a real example

Ties and actors' characteristics can change over time.



## Networks are dynamic by nature: a real example

Ties and actors' characteristics can change over time.



# Networks are dynamic by nature: a real example

Ties and actors' characteristics can change over time.



1. Social network dynamics can depend on actors' characteristics.

Selection process: relationship *partners* are selected according to their characteristics

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1. Social network dynamics can depend on actors' characteristics.

Selection process: relationship *partners* are selected according to their characteristics

### Example

Homophily: the formation of relations based on the similarity of two actors

### E.g. smoking behavior



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2. Changeable actors' characteristics can depend on the social network

 $\mathsf{E}.\mathsf{g}.:$  opinions, attitudes, intentions, etc. - we use the word behavior for all of these!

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Influence process: actors adjust their characteristics according to the characteristics of other actors to whom they are tied

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#### Example

Assimilation/contagion: connected actors become increasingly similar over time



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#### E.g. smoking behavior

Homophily and assimilation give rise to the same outcome (similarity of connected individuals)

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study of influence requires the consideration of selection and vice versa.

Fundamental question: is this similarity caused mainly by influence or mainly by selection?

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Extending the SAOM for the co-evolution of networks and behaviors

### Example

Similarity in smoking:

Selection: "a smoker may tend to have smoking friends because, once somebody is a smoker, he or she is likely to meet other smokers in smoking areas and thus has more opportunities to form friendship ties with them"

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### Example

Similarity in smoking:

Selection: "a smoker may tend to have smoking friends because, once somebody is a smoker, he or she is likely to meet other smokers in smoking areas and thus has more opportunities to form friendship ties with them"

Influence: "a smoker may have been the friendship with a smoker that made him or her start smoking in the first place"

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## Longitudinal network-behavior panel data

- 1. a network x represented by its adjacency matrix
- 2. a series of actors' attributes:
  - H constant covariates  $V_1, \cdots, V_H$
  - *L* behavior covariates  $Z_1(t), \dots, Z_L(t)$ Behavior variables are ordinal categorical variables.

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Longitudinal network-behavior panel data: networks and behaviors observed at  $M\geq 2$  time points  $t_1,\cdots,t_M$ 

$$(x,z)(t_1), (x,z)(t_2), \cdots, (x,z)(t_M)$$

and the constant covariates  $V_1, \cdots, V_H$ .

1. Distribution of the process.

Changes between observational time points are modeled according to a continuous-time Markov chain.

- State space  $\mathbb{C}$  : all the possible configurations arising from the combination of network and behaviors

$$|C| = 2^{n(n-1)} \times B^n$$

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where B is the number of categories for the behavior variable.

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#### 2. Opportunity to change.

At any given moment one probabilistically selected actor has the opportunity to change one of his outgoing tie or his behavior.



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The moments at which an actor has the opportunity for a tie change or a behavior change are modeled by two distinct rate functions.

#### 3. Absence of co-occurrence.

At each instant t, only one actor has the opportunity to change (one of his outgoing ties or his behavior)

4. Actor-oriented perspective.

Actors control their outgoing ties as well as their own behavior.

- the actor decide to change one of his outgoing ties or his behavior according to his position in the network, his attributes and the characteristics of the other actors

#### Aim: maximize a utility function

- two distinct objective functions: one for the network and one for the behavior change
- actors have complete knowledge about the network and the behaviors of all the the other actors
- the maximization is based on immediate returns and not on long-run rewarding (myopic actors)

# Model definition

According to the previous assumptions, the network-behavior co-evolution process is decomposed into a series of micro-steps:

- the opportunity of changing one network tie and the corresponding tie changed

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- the opportunity of changing a behavior and the corresponding unit changed in behavior

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- the opportunity of changing one network tie and the corresponding tie changed
- the opportunity of changing a behavior and the corresponding unit changed in behavior

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every micro-step requires the identification of a focal actor who gets the opportunity to make a change and the identification of the change outcome

	Occurrence	Preference
Network changes	Network rate function	Network objective function
Behavioral changes	Behavioral rate function	Behavioral objective function

# The rate functions

The frequency by which actors have the opportunity to make a change is modeled by the *rate functions*, one for each type of change.

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Why must we specify two different rate functions?

# The rate functions

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Practically always, one type of decision will be made more frequently than the other

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Why must we specify two different rate functions?

Practically always, one type of decision will be made more frequently than the other

### Example

In the joint study of friendship and smoking behavior at high school, we would expect more frequent changes in the network than in behavior

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#### Network rate function

 $T_i^{net}$  = the waiting time until *i* gets the opportunity to make a network change

 $T_i^{net} \sim Exp(\lambda_i^{net})$ 

### Behavior rate function

 $T_i^{beh}$  = the waiting time until *i* gets the opportunity to make a behavior change

 $T_i^{beh} \sim Exp(\lambda_i^{beh})$ 

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The waiting time until the occurrence of the next micro step of either kind by any actor is exponentially distributed with parameter:

$$\lambda_{tot} = \sum_{i} (\lambda_i^{net} + \lambda_i^{beh})$$

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#### Probabilities

 $P(i \text{ has the opportunity to change one of his tie}) = \frac{\lambda_i^{net}}{\lambda_{tot}}$  $P(i \text{ has the opportunity to change his behavior}) = \frac{\lambda_i^{beh}}{\lambda_{tot}}$ 

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## The rate functions (simplest specification)

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The waiting time until the occurrence of the next micro step of either kind by any actor is exponentially distributed with parameter:

$$\lambda_{tot} = n(\lambda^{net} + \lambda^{beh})$$

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### The rate functions (simplest specification)

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The waiting time until the occurrence of the next micro step of either kind by any actor is exponentially distributed with parameter:

$$\lambda_{tot} = n(\lambda^{net} + \lambda^{beh})$$

Probabilities

$$P(\text{network micro-step}) = \frac{n\lambda^{net}}{\lambda_{tot}}$$
  
 $P(\text{behavioral micro-step}) = \frac{n\lambda^{beh}}{\lambda_{tot}}$ 

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The probability of going from one state to another state of the co-evolution Markov chain is defined by the objective functions.

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Why must we specify two different objective functions?

The probability of going from one state to another state of the co-evolution Markov chain is defined by the objective functions.

 $\sum_{i=1}^{n}$  Why must we specify two different objective functions?

- The network objective function represents how likely it is for *i* to change one of its outgoing tie
- The behavioral objective function represents how likely it is for the actor *i* the current level of his behavior

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- The network objective function represents how likely it is for *i* to change one of its outgoing tie
- The behavioral objective function represents how likely it is for the actor i the current level of his behavior

#### Network objective function

$$f_i^{net}(\beta, x(i \rightsquigarrow j), z) = \sum_{k=1}^{K} \beta_k s_{ik}^{net}(x, z) + U_i(t, x, j)$$

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already described for the SAOM

Behavioral objective function

$$f_i^{beh}(\gamma, z(l \rightsquigarrow l'), x) = \sum_{w=1}^{W} \gamma_w s_{iw}^{beh}(x, z(l \rightsquigarrow l')) + \epsilon_i(t, z, l, l')$$

where

- $s_w^{beh}(x,z)$  are effects
- $\gamma_w$  are statistical parameters
- $\epsilon_i(t, z, l, l')$  is a random term

The probability that an actor *i* changes his own behavior by one unit is:

$$p_{ll'}(\gamma; z(l \rightsquigarrow l'); x) = \frac{\exp\left(\sum_{w=1}^{W} \gamma_w s_{lw}^{beh}(x, z(l \rightsquigarrow l'))\right)}{\sum_{l'' \in \{l+1, l-1, l\}} \exp\left(\sum_{w=1}^{W} \gamma_w s_{lw}^{beh}(x, z(l \rightsquigarrow l''))\right)}$$

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 $p_{II}$  is the probability that *i* does not change his behavior

### The specification of the behavioral objective function

- Basic shape effects

The linear shape effect  $s_{i\_linear}^{beh}(x,z)$  and the quadratic shape effect  $s_{i\_quadratic}^{beh}(x,z)$  are defined by

$$s_{i\_linear}^{beh}(x,z) = z_i$$
  $s_{i\_quadratic}^{beh}(x,z) = z_i^2$ 

The basic shape effects must be always included in the model specification.

#### The specification of the behavioral objective function

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### The specification of the behavioral objective function

- Classical influence effects
  - 1. The average similarity effect  $s_{i_avsim}^{beh}(x,z)$  expressing the preference of actors to be similar in behavior to their alters, in such a way that the total influence of the alters is the same regardless ego's outdegree

$$s_{i\_avsim}^{beh}(x,z) = \frac{1}{x_{i+}} \sum_{j=1}^{n} x_{ij}(sim_z(ij) - sim_z)$$

where

$$sim_z(ij) = 1 - rac{\left|z_i - z_j\right|}{R_z}$$

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 $R_z$  is the range of the behavior z and  $sim_z$  is the mean similarity value.

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- Classical influence effects
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 $R_z$  is the range of the behavior z and  $sim_z$  is the mean similarity value.

 The total similarity effect s<sup>beh</sup><sub>i-totsim</sub>(x, z), expressing the preference of actors to be similar in behavior to their alters, in such a way that the total influence of the alters is proportional to the number of alters:

$$s_{i\_totsim}^{beh}(x,z) = \sum_{j=1}^{n} x_{ij}(sim_z(ij) - sim_z)$$

### The specification of the behavioral objective function

- Position-dependent influence effects
   Network position could also have an effect on the behavior of dynamics.
  - 1. outdegree effect

$$s_{i\_out}^{beh}(x,z) = z_i \sum_{j=1}^n x_{ij}$$

2. indegree effect

$$s_{i\_ind}^{beh}(x,z) = z_i \sum_{j=1}^n x_{ji}$$

- Effects of other actor variables.

For each actor's attribute a main effect on the behavior can be included in the model.

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The algorithm consists in reproducing a possible series of micro-steps between two observation moments  $t_0$  and  $t_1$  according to fixed parameter value.

- 1. Set the time t = 0,  $x = x(t_0)$  and  $z = z(t_0)$
- 2. Generate  $T^{\mathit{net}}$  according to an exponential distribution with parameter  $\lambda^{\mathit{net}}$
- 3. Generate  ${\cal T}^{beh}$  according to an exponential distribution with parameter  $\lambda^{beh}$

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4. If min  $\{T^{net}, T^{beh}\} = T^{net}$  a **network micro-step** is implemented:

- Select the actor  $i \in \mathcal{N}$ , who makes the changes, with probability

 $P(i \text{ has the opportunity to change } | \text{ network micro-steps}) = \frac{\lambda^{net}}{\lambda_{tot}}$ 

- Select the actor  $j \in \mathcal{N}$ , to whom *i* changes his outgoing tie, with probability:

$$p_{ij} = \frac{\exp\left(\sum_{k=1}^{K} \beta_k s_{ik}(x(i \rightsquigarrow j), z)\right)}{\sum_{h=1}^{n} \exp\left(\sum_{k=1}^{K} \beta_k s_{ik}(x(i \rightsquigarrow h), z)\right)}$$

- If 
$$i \neq j$$
 then  $x = x(i \rightsquigarrow j)$ . If  $i = j$  then  $x = x$   
- Set  $t = t + T^{net}$ 

4. If min  $\{T^{net}, T^{beh}\} = T^{beh}$  a behavioral micro-step is implemented:

- Select the actor  $i \in \mathcal{N}$ , who makes the changes, with probability

 $P(i \text{ has the opportunity to change}|\text{behavioral micro-steps}) = \frac{\lambda^{beh}}{\lambda_{tot}}$ 

- Determine the behavioral change  $l' \in \{l+1, l-1, l\}$  with probability:

$$p_{II'}(\gamma; z(l \rightsquigarrow l'); x) = \frac{\exp\left(\sum_{w=1}^{W} \gamma_w s_{iw}^{beh}(x, z(l \rightsquigarrow l'))\right)}{\sum_{l'' \in \{l+1, l-1, l\}} \exp\left(\sum_{w=1}^{W} \gamma_w s_{iw}^{beh}(x, z(l \rightsquigarrow l''))\right)}$$

- If 
$$l \neq l'$$
 then  $z = z(l \rightsquigarrow l')$ . If  $l = l'$  then  $z = z$   
- Set  $t = t + T^{beh}$ 

5. Repeat step 2. to 4. until the stopping criterion is satisfied.

1. Unconditional simulation:

the simulations in each time period carry on until a predetermined time length has elapsed (usually until t = 1).

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1. Unconditional simulation:

the simulations in each time period carry on until a predetermined time length has elapsed (usually until t = 1).

- 2. Conditional simulation on the observed number of changes:
  - simulations run on until the number of different entries between  $x(t_0)$  and the simulated network  $x^{sim}(t_1)$  is equal to the number of entries that differ between  $x(t_0)$  and  $x(t_1)$

$$\sum_{\substack{i,j=1\\ i\neq j}}^{n} \left| X_{ij}^{obs}(t_1) - X_{ij}(t_0) \right| = \sum_{\substack{i,j=1\\ i\neq j}}^{n} \left| X_{ij}^{sim}(t_1) - X_{ij}(t_0) \right|$$

- simulations run on until the number of different entries between  $z(t_0)$  and the simulated behavior  $z^{sim}(t_1)$  is equal to the number of entries that differ between  $z(t_0)$  and  $z(t_1)$ 

$$\sum_{i=1}^{n} \left| z_{i}^{obs}(t_{1}) - z_{i}(t_{0}) \right| = \sum^{n} \left| z_{i}^{sim}(t_{1}) - z_{i}(t_{0}) \right|$$

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Aim: estimate  $\theta$ , the 2(M-1)+K+W dimensional vector of parameters of the co-evolution model

### Statistics:

- Network rate parameters for the period m

$$s_{\lambda_m}^{net}(X(t_m), X(t_{m-1})|X(t_{m-1}) = x(t_{m-1})) = \sum_{\substack{i,j=1\\ i\neq j}}^n |X_{ij}(t_m) - X_{ij}(t_{m-1})|$$

- Behavior rate parameters for the period m

$$s_{\lambda_m}^{beh}(Z(t_m), Z(t_{m-1})|Z(t_{m-1}) = z(t_{m-1})) = \sum_{i=1}^n |Z_i(t_m) - Z_i(t_{m-1})|$$

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Aim: estimate  $\theta$ , the 2(M-1)+K+W dimensional vector of parameters of the co-evolution model

#### Statistics:

- Network objective function effects

$$\sum_{m=1}^{M-1} s_{mk}^{net}(X(t_m)|X(t_{m-1}) = x(t_{m-1})) = \sum_{m=1}^{M-1} s_{mk}^{net}(X(t_m), X(t_{m-1}))$$

- Behavior objective function effects

$$\sum_{m=1}^{M-1} s_{mw}^{beh}(X(t_m)|X(t_{m-1}) = x(t_{m-1})) = \sum_{m=1}^{M-1} s_{mw}^{beh}(X(t_m), X(t_{m-1}))$$

Consequently the MoM estimator for  $\theta$  is provided by the solution of the system of equations:

$$\begin{cases} E_{\theta} \left[ s_{\lambda_{m}} (X(t_{m}), X(t_{m-1}) | X(t_{m-1}) = x(t_{m-1})) \right] = s_{\lambda_{m}} (x(t_{m}), x(t_{m-1})) \\ E_{\theta} \left[ s_{\lambda_{m}} (Z(t_{m}), Z(t_{m-1}) | Z(t_{m-1}) = z(t_{m-1})) \right] = s_{\lambda_{m}} (z(t_{m}), z(t_{m-1})) \\ E_{\theta} \left[ \sum_{m=1}^{M-1} s_{mk}^{net} (X(t_{m}) | X(t_{m-1}) = x(t_{m-1})) \right] = \sum_{m=1}^{M-1} s_{mk}^{net} (x(t_{m}), x(t_{m-1})) \\ E_{\theta} \left[ \sum_{m=1}^{M-1} s_{mw}^{beh} (X(t_{m}) | X(t_{m-1}) = x(t_{m-1})) \right] = \sum_{m=1}^{M-1} s_{mw}^{beh} (x(t_{m}), x(t_{m-1})) \end{cases}$$

### Remarks

- The system of equation cannot be solved analytically  $\Longrightarrow$  Robbins-Monro algorithm
- The Maximum likelihood estimation is under construction. At the moment is too slow

## Example

Example data: excerpt from the "Teenage Friends and Lifestyle Study" data set

We will use the SAOM for the co-evolution of networks and behaviors to disentangle influence and selection processes.

- 1. Do pupils select friends based on similar smoking behavior?
- 2. Are pupils influenced by friends to adjust to their smoking behavior?

Dependent variables: friendship networks and smoking behavior Covariate: gender To find out whether it makes sense to analyze the data with a co-evolution model one should check whether:

1. the data are sufficiently informative to allow for identification of effects

$$J = \frac{N_{11}}{N_{11} + N_{01} + N_{10}} > 0.3 \qquad \qquad Jaccard \ index$$

Tie changes between subsequent observations: periods 0 =>  $1 \implies 0$ Distance Jaccard Missing 0 => 0 1 1 => 11 ==> 2 15827 237 240 208 477 0.304 0 (0%) 3 2 ==> 15839 228 209 236 437 0.351 0 (0%)

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### Precondition of the analysis

2. there is interdependence between networks and behavioral variables

$$I = -\frac{\overline{d}(\delta)}{(n-1)\overline{d}}$$
 Moran index

where

$$\overline{d}(\delta) = rac{\sum\limits_{ij} x_{ij}(z_i - \overline{z})(z_j - \overline{z})}{\sum\limits_{ij} x_{ij}}$$

is the mean of the cross products of the centered behavioral variable for connected actors

$$\overline{d} = \frac{\sum_{ij} (z_i - \overline{z})(z_j - \overline{z})}{n(n-1)}$$

is the overall mean of the cross products of the centered behavioral variable for all the possible pairs of actors in the network

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$$\frac{\sum_{ij} x_{ij}(z_i - \overline{z})(z_j - \overline{z})}{\sum_{ij} x_{ij}} \qquad \frac{\sum_{ij} (z_i - \overline{z})(z_j - \overline{z})}{n(n-1)}$$

If two connected actors i and j are similar in their behaviors we expect that

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- $z_i > \overline{z}$  and  $z_j > \overline{z}$
- $z_i < \overline{z}$  and  $z_j < \overline{z}$



$$\frac{\sum\limits_{ij} x_{ij}(z_i - \overline{z})(z_j - \overline{z})}{\sum\limits_{ij} x_{ij}} > \frac{\sum\limits_{ij} (z_i - \overline{z})(z_j - \overline{z})}{n(n-1)}$$

If two connected actors i and j are similar in their behaviors we expect that

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- 
$$z_i > \overline{z}$$
 and  $z_j > \overline{z}$ 

-  $z_i < \overline{z}$  and  $z_j < \overline{z}$ 



$$\frac{\sum\limits_{ij} x_{ij}(z_i - \overline{z})(z_j - \overline{z})}{\sum\limits_{ij} x_{ij}} > \frac{\sum\limits_{ij} (z_i - \overline{z})(z_j - \overline{z})}{n(n-1)}$$

If two connected actors i and j are similar in their behaviors we expect that

- 
$$z_i > \overline{z}$$
 and  $z_j > \overline{z}$ 

- 
$$z_i < \overline{z}$$
 and  $z_j < \overline{z}$ 

 $\underset{\text{Positive interdependence}}{\Downarrow}$ 

0 < l < 1

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$$\frac{\sum_{ij} x_{ij}(z_i-z)(z_j-z)}{\sum_{ij} x_{ij}} \qquad \frac{\sum_{ij} (z_i-z)(z_j-z)}{n(n-1)}$$

If two connected actors i and j are extremely different in their behaviors we expect that

- 
$$z_i > \overline{z}$$
 and  $z_j < \overline{z}$ 

- 
$$z_i < \overline{z}$$
 and  $z_j > \overline{z}$ 





$$\frac{\sum\limits_{ij} x_{ij}(z_i - \overline{z})(z_j - \overline{z})}{\sum\limits_{ij} x_{ij}} < \frac{\sum\limits_{ij} (z_i - \overline{z})(z_j - \overline{z})}{n(n-1)}$$

If two connected actors i and j are extremely different in their behaviors we expect that

- $z_i > \overline{z}$  and  $z_j < \overline{z}$
- $z_i < \overline{z}$  and  $z_j > \overline{z}$



$$\frac{\sum\limits_{ij} x_{ij}(z_i - \overline{z})(z_j - \overline{z})}{\sum\limits_{ij} x_{ij}} < \frac{\sum\limits_{ij} (z_i - \overline{z})(z_j - \overline{z})}{n(n-1)}$$

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 $\underset{\text{Negative interdependence}}{\Downarrow}$ 

-1 < I < 0

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Precondition of the analysis: understanding the Moran index

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#### Precondition of the analysis: understanding the Moran index

lf:



$$\frac{\sum_{ij} x_{ij}(z_i - \overline{z})(z_j - \overline{z})}{\sum_{ij} x_{ij}} \approx \frac{\sum_{ij} (z_i - \overline{z})(z_j - \overline{z})}{n(n-1)}$$

the *local* distribution of the behavioral variable follows the *global* distribution of the behavioral variable

Absence of interdependence

$$I \approx -\frac{1}{n-1}$$

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#### Precondition of the analysis

#### In theory $-1 \leq l \leq +1$

- values close to zero indicates independence between networks and behaviors (i.e.absence of interdependence)
- value +1 indicates perfect identity of the behaviors of two friends (i.e. very strong positive interdependence)
- value -1 indicates perfect complementarity of the behaviors of two friends (i.e. very strong negative interdependence).

#### Precondition of the analysis

In theory  $-1 \le l \le +1$ 

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The computation of the Moran index for the friendship networks and smoking behaviors leads to

0.244 0.258 0.341

Conclusion: there is considerable dependence between networks and behaviors.

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	Estimates	s.e.	t-score
Network Dynamics			
constant friendship rate (period 1)	8.6287	(0.6666)	
constant friendship rate (period 2)	7.2489	(0.5466)	
outdegree (density)	-2.4084	(0.0407)	-59.12676
reciprocity	2.7024	( 0.0823 )	32.8337
Behavior Dynamics			
rate smokebeh (period 1)	3.8922	( 1.9689 )	
rate smokebeh (period 2)	4.4813	(2.3679)	
behavior smokebeh linear shap	-3.5464	(0.4394)	-8.0712
behavior smokebeh quadratic shape	2.8464	(0.3628)	7.8447

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rate smokebeh (period 2)	4 4813	(2.3679)	
rate smokebell (period 2)	1.1010	(2.0015)	
behavior smokebeh linear shap	-3.5464	(0.4394)	-8.0712
behavior smokebeh quadratic shape	2.8464	(0.3628)	7.8447

#### Network rate parameters:

- about 9 opportunities for a network change in the first period
- about 7 opportunities for a network change in the second period

	Estimates	s.e.	t-score
Network Dynamics			
constant friendship rate (period 1)	8.6287	(0.6666)	
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Network objective function parameters:

- outdegree parameter: the observed networks have low density
- reciprocity parameter: strong tendency towards reciprocated ties

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#### Behavioral rate parameters:

- about 4 opportunities for a behavioral change in the first period
- about 4 opportunities for a behavioral change in the second period

	Estimates	s.e.	t-score
Network Dynamics			
constant friendship rate (period 1)	8.6287	(0.6666)	
constant friendship rate (period 2)	7.2489	( 0.5466 )	
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Behavioral objective function parameters: express the attractiveness of different behavioral levels taking into account the current structure of the network and the behavior of the other actors

- Smoking behavior: coded with 1 for "no", 2 for "occasional", and 3 for "regular" smokers.

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- The smoking covariate is centered:  $\overline{z} = 1.377$  is the mean of the covariate

$$z_i - \overline{z} = \left\{ \begin{array}{ll} -0.377 & \mbox{ for no smokers} \\ 0.623 & \mbox{ for occasional smokers} \\ 1.623 & \mbox{ for regular smokers} \end{array} \right.$$

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- The smoking covariate is centered:  $\overline{z} = 1.377$  is the mean of the covariate

$$z_i - \overline{z} = \begin{cases} -0.377 & \text{for no smokers} \\ 0.623 & \text{for occasional smokers} \\ 1.623 & \text{for regular smokers} \end{cases}$$

- The contribution to the behavioral objective function is

$$\gamma_{linear}(z_i - \overline{z}) + \gamma_{quadratic}(z_i - \overline{z})^2 =$$
$$= -3.5464_{linear}(z_i - \overline{z}) + 2.8464_{quadratic}(z_i - \overline{z})^2$$

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U-shaped changes in the behavior are drawn to the extreme of the range

The baseline model does not provide any information about selection and influence processes:

- the network dynamics are explained by the preference towards creating and reciprocating ties
- the behavior dynamic are described only by the distribution of the behavior in the population

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- the network dynamics are explained by the preference towards creating and reciprocating ties
- the behavior dynamic are described only by the distribution of the behavior in the population

If we want to disentangle the selection and influence effects we should include in the objective functions specification:

- the effects that capture the dependence of social network dynamics on actor's characteristic
- the effects that capture the dependence of behavior dynamics on social network

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# Effects that capture the dependence of social network dynamics on actor's characteristic

- pupils prefer to establish friendship relations with others that are similar to themselves

## Effects that capture the dependence of social network dynamics on actor's characteristic

- pupils prefer to establish friendship relations with others that are similar to themselves  $\rightarrow$  covariate similarity



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## Effects that capture the dependence of social network dynamics on actor's characteristic

- pupils prefer to establish friendship relations with others that are similar to themselves  $\rightarrow$  covariate similarity



This effect must be controlled for the sender and receiver effects of the covariate.

- Covariate ego effect



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# Effects that capture the dependence of behavior dynamics on social network

- pupils tend to adjust their smoking behavior according to the behaviors of their friends

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## Effects that capture the dependence of behavior dynamics on social network

- pupils tend to adjust their smoking behavior according to the behaviors of their friends  $\rightarrow$  average similarity effect



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# Effects that capture the dependence of behavior dynamics on social network

- pupils tend to adjust their smoking behavior according to the behaviors of their friends  $\rightarrow$  average similarity effect



This effect must be controlled for the indegree and the outdegree effects

- Indegree effect



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- Outdegree effect

	Estimates	s.e.	t-score
Network Dynamics			
constant friendship rate (period 1)	10.7166	(1.4036)	
constant friendship rate (period 2)	9.0005	(0.7709)	
outdegree (density)	-2.8435	(0.0572)	-49.6776
reciprocity	1.9683	(0.0933)	21.1077
transitive triplets	0.4447	(0.0322)	13.7964
sex ego	0.1612	(0.1206)	1.3368
sex alter	-0.1476	(0.1064)	-1.3871
sex similarity	0.9104	(0.0882)	10.3244
smoke ego	0.0665	(0.0846)	0.7857
smoke alter	0.1121	(0.0761)	1.4719
smokebeh similarity	0.5114	(0.1735)	2.9479

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Network objective function parameters: tendency towards reciprocity, transitivity and homophily with respect to gender

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Network objective function parameters: pupils selected others with similar smoking behavior as friends  $\rightarrow$  evidence for selection process

	Estimates	s.e.	t-score
Behavior Dynamics			
rate smokebeh (period 1)	3.9041	(1.7402)	
rate smokebeh (period 2)	3.8059	(1.4323)	
		. ,	
behavior smokebeh linear shap	-3.3573	(0.5678)	-5.9129
behavior smokebeh quadratic shape	2.8406	(0.4125)	6.8864
behavior smokebeh indegree	0.1711	(0.1812)	0.9444
behavior smokebeh outdegree	0.0128	(0.1926)	0.0662
behavior smokebeh average similarity	3.4361	(1.4170)	2.4250

Behavioral objective function parameters: U-shaped distribution of the smoking behavior

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Behavioral objective function parameters: indegree and outdegree effects are not significant

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Behavior Dynamics			
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behavior smokebeh average similarity	3.4361	(1.4170)	2.4250

Behavioral objective function parameters: pupils are influenced by the smoking behavior of the others  $\to$  evidence for influence process

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The contribution to the behavioral objective function is given by:

$$\gamma_{\textit{linear}}(z_i - \overline{z}) + \gamma_{\textit{quadratic}}(z_i - \overline{z})^2 + \gamma_{\textit{avsim}} \frac{1}{x_{i+}} \sum_{j=1}^n x_{ij}(\textit{sim}_z(ij) - \textit{sim}_z) =$$

$$= -3.3573_{linear}(z_i - \overline{z}) + 2.8406_{quadratic}(z_i - \overline{z})^2 + 3.4361\frac{1}{x_{i+}}\sum_{j=1}^n x_{ij}(sim_z(ij) - 0.7415)$$

Since the behavioral objective function depends not generally on the average behavior of the actor's friends, here we present a table only for the special case of actors all whose friend have the same behavior  $z_i$ .

z <sub>j</sub> / z <sub>i</sub>	no	occasional	regular
1	2.56	-1.82	-0.51
2	0.84	-0.10	1.20
3	-0.88	-1.82	2.92

- The row maximum is assumed at the diagonal for the non-smokers and for the regular smokers  $\rightarrow$  the focal actor prefers to have the same behavior as all these friends.
- In the case where the friends do not smoke at all, the preference toward imitating their behavior is less strong than in the case where all the friends smoke a lot.

#### Recent, current and near future

- Distinction among creating and deleting ties
- Estimation procedures (MLE and bayesian estimation)

- Goodness of fit of the model
- Model selection
- Non-directed relations
- Time-heterogeneity tests
- ...